



The Kalman Filter

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Overview

- ▶ Recursive Least Squares Revisited
- ▶ Discrete Time Kalman Filter
- ▶ Continuous Time Kalman Filter
- ▶ Extended Kalman Filter

Comments

Recursive Least Squares (RLS) Revisited

For linear models $y_k = \phi_k^\top \theta + \epsilon_k$ and i.i.d. Gaussian noise, we had the following recursion.

- ▶ start with some a-priori knowledge on θ in form of a mean $\hat{\theta}_0$ and inverse covariance Q_0 . Start loop at $k = 1$
- ▶ at time k , when measurement y_k is known, compute new inverse covariance k as

$$Q_k = Q_{k-1} + \phi_k \phi_k^\top$$

(adding measurements increases "information")

- ▶ compute new estimate for the mean ("innovation update")

$$\hat{\theta}_k = \hat{\theta}_{k-1} + \underbrace{Q_k^{-1} \phi_k (y_k - \phi_k^\top \hat{\theta}_{k-1})}_{\text{"innovation"}}$$

$$y(1) \equiv y_1$$

$$\hat{\theta}_{k-1}$$

$$\{y_1, y_2, y_3, \dots, y_{k-1}, y_k\} \rightarrow \hat{\theta}_k$$

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$$Q_k = \underset{\alpha}{Q_{k-1}} + \phi_k \phi_k^\top \quad \text{rec}(0,1)$$

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- ▶ delivers recursive solution to LLS problem

$$\hat{\theta}_k = \arg \min_{\theta} \quad \underbrace{(\theta - \hat{\theta}_0)^\top Q_0 (\theta - \hat{\theta}_0)} + \sum_{i=1}^k (y_i - \phi_i^\top \theta)^2$$

Comments

Recursive Least Squares for State Estimation (1)

- Assume now that we have a known deterministic linear system

$$x_i = A_{i-1}x_{i-1}$$

with linear measurement equation (for $i = 1, \dots, k$)

$$y_i = C_i x_i + v_i$$

(where v_i is i.i.d. ^{Gaussian} zero mean noise), and initial knowledge on initial state in form of mean \hat{x}_0 and inverse covariance Q_0

- Clearly, $x_k = A_{k-1} \cdots A_0 x_0$ and thus

$$y_k = C_k A_{k-1} \cdots A_0 x_0 + v_k$$

- Using $\theta \equiv x_0$, $\epsilon_k \equiv v_k$ and $\phi_k^\top = C_k A_{k-1} \cdots A_0$, this can be cast into the standard RLS framework from before

Comments

Recursive Least Squares for State Estimation (2)

- ▶ With $\theta \equiv x_0$ and

$$\phi_k^\top = C_k A_{k-1} \cdots A_0$$

our recursion becomes:

$$Q_k = Q_{k-1} + (C_k A_{k-1} \cdots A_0)^\top C_k A_{k-1} \cdots A_0$$

and

$$\hat{\theta}_k = \hat{\theta}_{k-1} + Q_k^{-1} (C_k A_{k-1} \cdots A_0)^\top (y_k - C_k A_{k-1} \cdots A_0 \hat{\theta}_{k-1})$$

- ▶ often, we are most interested in current state x_k . Let us denote its estimate given the data (y_1, \dots, y_m) by $\hat{x}_{[k|m]}$. Clearly,

$$\hat{x}_{[k|m]} = A_{k-1} \cdots A_0 \hat{\theta}_m$$

and

$$\underbrace{\text{Cov}\{x_{[k|m]}\}}_{=: P_{[k|m]}} = A_{k-1} \cdots A_0 \cdot Q_m^{-1} \cdot A_0^\top \cdots A_{k-1}^\top$$

Comments

Recursive Least Squares for State Estimation (3)

- ▶ the innovation update of the mean then simplifies to

$$\hat{x}_{[k|k]} = \hat{x}_{[k|k-1]} + \underline{P_{[k|k]}} \cdot \underline{C_k^\top (y_k - C_k \hat{x}_{[k|k-1]})} \quad 3$$

- ▶ the covariance update becomes (assuming for simplicity invertibility of $A_{k-1} \cdots A_0$)

$$P_{[k|k]}^{-1} = P_{[k|k-1]}^{-1} + C_k^\top C_k \quad 2$$

- ▶ for the propagation in time (prediction), we can use

$$\underline{\hat{x}_{[k|k-1]}} = A_{k-1} \cdot \underline{\hat{x}_{[k-1|k-1]}} \quad \left. \begin{array}{l} \\ \end{array} \right\} 1$$

and

$$P_{[k|k-1]} = A_{k-1} \cdot \underline{P_{[k-1|k-1]}} \cdot A_{k-1}^\top$$

Comments

Summary of Recursive Least Squares

Two steps to compute estimates $\hat{x}_{[k|k]}$ and covariances $P_{[k|k]}$:

- Prediction Step (before measurement):

$$\hat{x}_{[k|k-1]} = A_{k-1} \cdot \hat{x}_{[k-1|k-1]} \quad (1)$$

$$P_{[k|k-1]} = A_{k-1} \cdot P_{[k-1|k-1]} \cdot A_{k-1}^\top \quad (2)$$

- Innovation Update Step (after measurement):

$$P_{[k|k]} = \left(P_{[k|k-1]}^{-1} + C_k^\top C_k \right)^{-1} \quad (3)$$

$$\hat{x}_{[k|k]} = \hat{x}_{[k|k-1]} + P_{[k|k]} \cdot C_k^\top (y_k - C_k \hat{x}_{[k|k-1]}) \quad (4)$$

- Can interpret $\hat{x}_{[k|k-1]}$ and $P_{[k|k-1]}$ as a-priori information on x_k , based on prediction model $x_k = A_{k-1}x_{k-1}$

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- Could we also incorporate i.i.d. state noise w_{k-1} with zero mean and covariance W_{k-1} ?

Summary of Recursive Least Squares (and Extension)

Two steps to compute estimates $\hat{x}_{[k|k]}$ and covariances $P_{[k|k]}$:

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- Innovation Update Step (after measurement):

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Comments

The Kalman Filter

We have basically obtained the Kalman Filter. More generally, can assume measurement noise v_k with covariance V , such that measurements are weighted with V^{-1} .

- ▶ Full model: $x_{k+1} = A_k x_k + w_k$ and $y_k = C_k x_k + v_k$ with i.i.d. zero mean noises with covariances W , V .
- ▶ The steps of the Kalman Filter are:
 - ▶ Prediction Step:

$$\hat{x}_{[k|k-1]} = A_{k-1} \cdot \hat{x}_{[k-1|k-1]} \quad (5)$$

$$P_{[k|k-1]} = A_{k-1} \cdot P_{[k-1|k-1]} \cdot A_{k-1}^\top + W_{k-1} \quad (6)$$

- ▶ Innovation Update Step:

$$P_{[k|k]} = \left(P_{[k|k-1]}^{-1} + C_k^\top \underline{V^{-1}} C_k \right)^{-1} \quad (7)$$

$$\hat{x}_{[k|k]} = \hat{x}_{[k|k-1]} + P_{[k|k]} \cdot C_k^\top \underline{V^{-1}} (y_k - C_k \hat{x}_{[k|k-1]}) \quad (8)$$



Comments

Optimization Problem solved by Kalman Filter

$$\begin{aligned} \text{Minimize}_{x_0, \dots, x_N} & (x_0 - \hat{x}_0)^T P_0^{-1} (x_0 - \hat{x}_0) \\ & + \sum_{i=1}^N (y_i - C_i x_i)^T V^{-1} (y_i - C_i x_i) \\ & + \sum_{k=0}^{N-1} (x_{k+1} - A_k x_k)^T W^{-1} (x_{k+1} - A_k x_k) \end{aligned}$$

$$x_0^*, \dots, x_{N-1}^* \rightarrow \boxed{x_N^*} \equiv x_{[1:N]}$$

Comments

The Kalman Filter in Continuous Time (1)

- ▶ Continuous time model with noises w^c , v^c :

$$\dot{x}(t) = A^c(t)x(t) + w^c(t)$$

$$y(t) = C^c(t)x(t) + v^c(t)$$

- ▶ Note that $w^c(t)$ has unit $[x]/[t]$ and $v^c(t)$ has unit $[y]$
- ▶ Assume white noises:
 - ▶ $\text{Cov}(w^c(t_1), w^c(t_2)) = \delta(t_1 - t_2) \cdot W^c$ where $[W^c] = [x]^2/[t]$
 - ▶ $\text{Cov}(v^c(t_1), v^c(t_2)) = \delta(t_1 - t_2) \cdot V^c$ where $[V^c] = [y]^2 \cdot [t]$
- ▶ Transfer to discrete time: use small time step Δt and time points $t_k = k \cdot \Delta t$, identify $x_k = x(t_k)$.
- ▶ Identify $w_k = \int_{t_k}^{t_{k+1}} w^c(t) dt$. Random walk theory says: $W = \Delta t \cdot W^c$ (the longer we wait, the more uncertain we become)
- ▶ Identify $C_k = C(t_k)$, $y_k = \frac{1}{\Delta t} \int_{t_k}^{t_{k+1}} y(t) dt$, and thus $v_k = \frac{1}{\Delta t} \int_{t_k}^{t_{k+1}} v^c(t) dt$. Due to averaging, covariance matrix shrinks with longer time intervals, i.e. $V = V^c / \Delta t$ (we give more weight to each measurement when taking fewer measurements).

Comments

The Kalman Filter in Continuous Time (2)

- Discrete time model:

$$x_{k+1} = \underbrace{[I + \Delta t \cdot A^c(t_k)]}_{=: A_k} x_k + w_k$$
$$y_k = \underbrace{C^c(t_k)}_{=: C_k} x_k + v_k$$

- Covariances:

$$\text{Cov}(w_k) = W = \Delta t \cdot W^c \text{ and}$$

$$\text{Cov}(v_k) = V = \Delta t^{-1} \cdot V^c$$

Comments

The Kalman Filter in Continuous Time (3)

- Prediction step (up to first order):

$$\begin{aligned}\hat{x}_{[k+1|k]} &= \hat{x}_{[k|k]} + \Delta t \cdot A^c(t_k) \hat{x}_{[k|k]} \\ P_{[k+1|k]} &= P_{[k|k]} + \Delta t \left[A^c(t_k) P_{[k|k]} + P_{[k|k]} A^c(t_k)^\top + W^c \right]\end{aligned}$$

- Innovation Update Step:

$$\begin{aligned}P_{[k|k]} &= \left(P_{[k|k-1]}^{-1} + \Delta t \cdot C_k^\top (V^c)^{-1} C_k \right)^{-1} \\ \hat{x}_{[k|k]} &= \hat{x}_{[k|k-1]} + \Delta t \cdot P_{[k|k]} \cdot C_k^\top (V^c)^{-1} (y_k - C_k \hat{x}_{[k|k-1]})\end{aligned} \tag{9}$$

- By Taylor expansion, Eq. (9) becomes

$$P_{[k|k]} = P_{[k|k-1]} - \Delta t \cdot P_{[k|k-1]} C_k^\top (V^c)^{-1} C_k P_{[k|k-1]}$$

Comments

The Kalman Filter in Continuous Time (4)

- ▶ Adding both steps together, identifying $\hat{x}(t_k) \equiv \hat{x}_{[k+1|k]}$ and $P(t_k) \equiv P_{[k+1|k]}$, and taking the limit $\Delta t \rightarrow 0$ yields the two differential equations:

$$\begin{aligned}\dot{\hat{x}}(t) &= A^c(t_k)\hat{x}(t) + P(t)C^c(t)^\top (V^c)^{-1} (y(t) - C^c(t)\hat{x}(t)) \\ \dot{P}(t) &= A^c(t)P(t) + P(t)A^c(t)^\top \\ &\quad + W^c - P(t)C^c(t)^\top (V^c)^{-1} C^c(t)P(t)\end{aligned}$$

- ▶ These continuous time equations are called the “Kalman-Bucy-Filter”

Comments

The Extended Kalman Filter in Continuous Time

Comments

The Extended Kalman Filter in Discrete Time

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